

データ解析(第13回)

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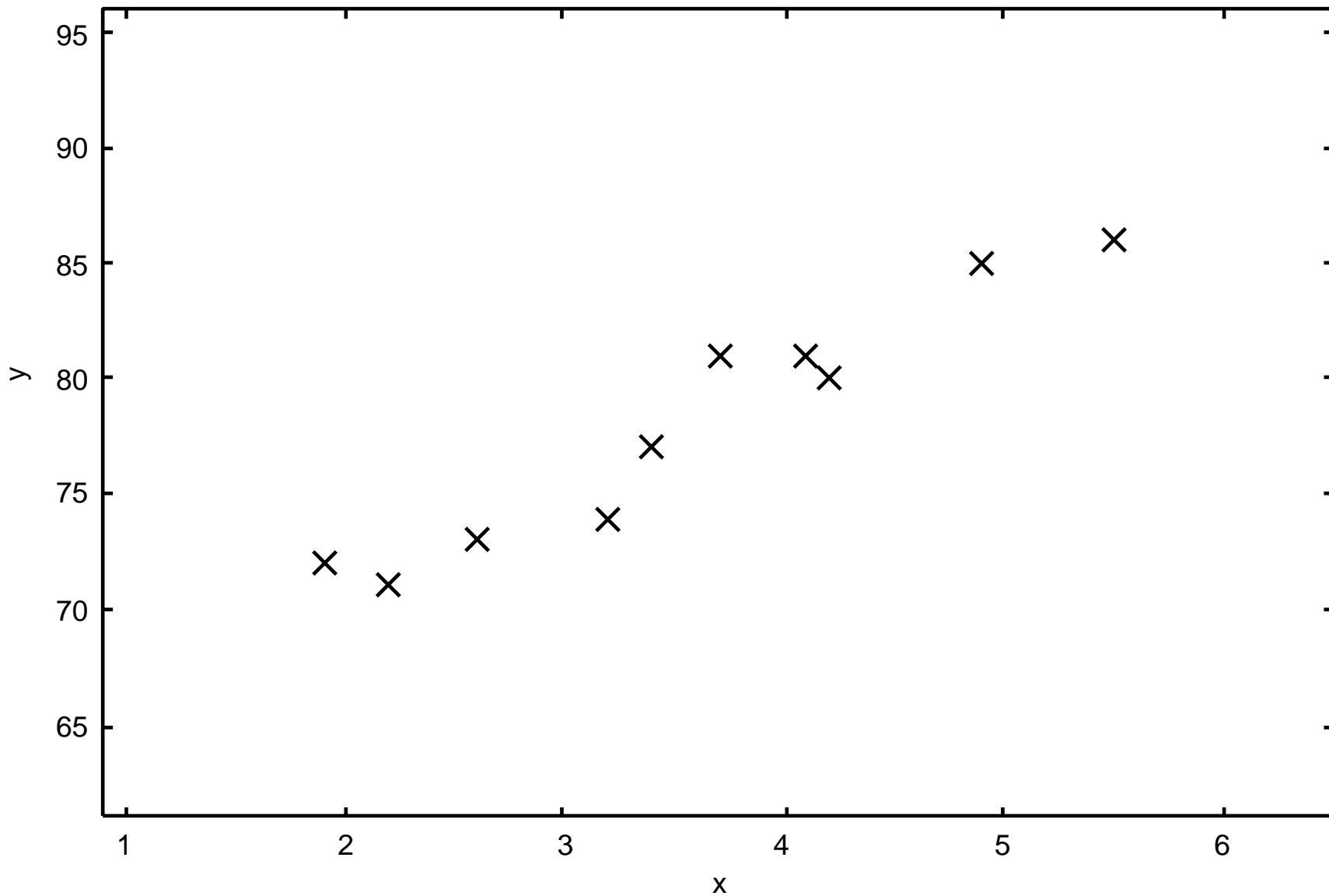
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第5章 重回帰分析

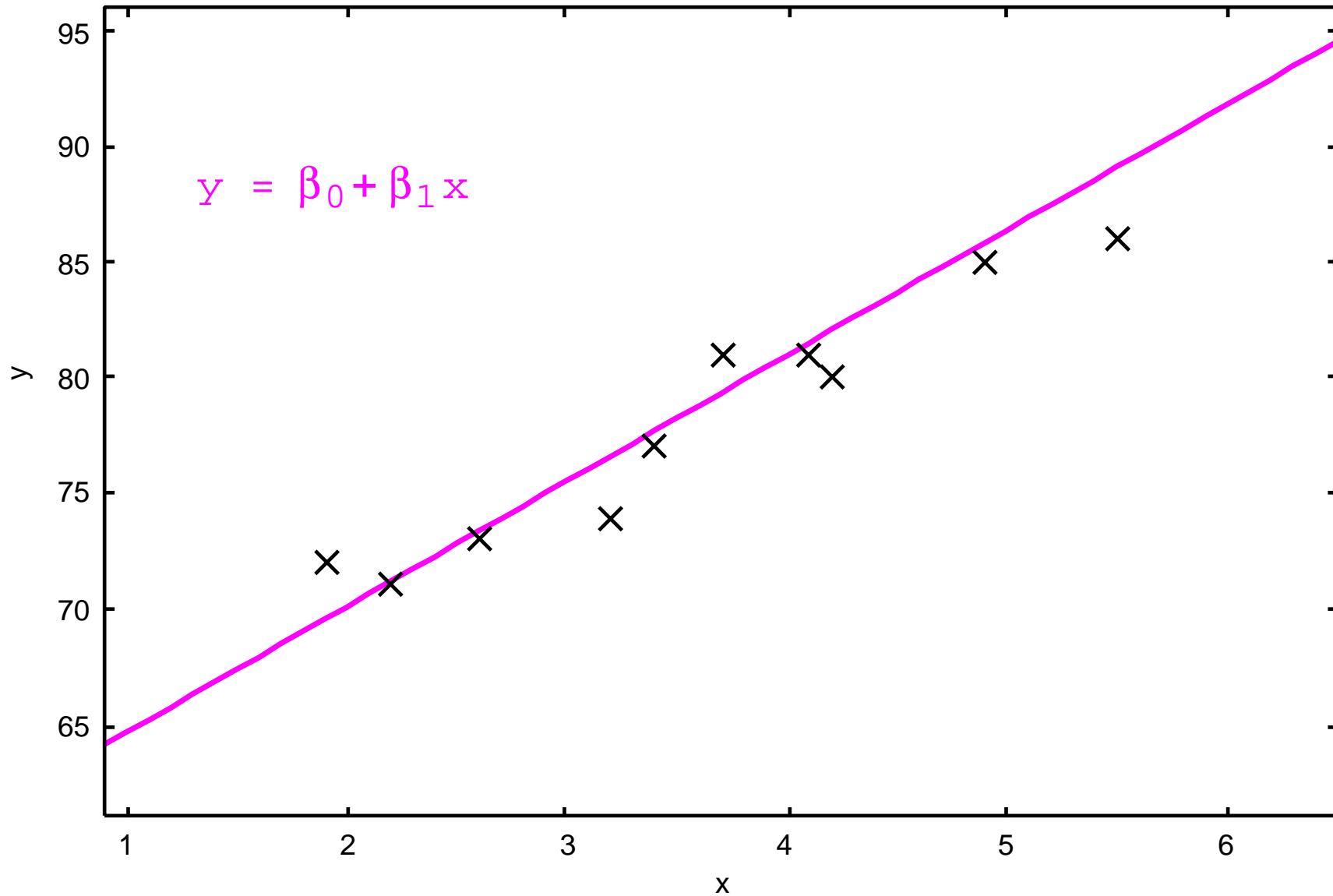
データ (x_i, y_i)

No.	説明変数	目的変数
	x	y
1	x_1	y_1
2	x_2	y_2
.	.	.
.	.	.
i	x_i	y_i
.	.	.
.	.	.
n	x_n	y_n

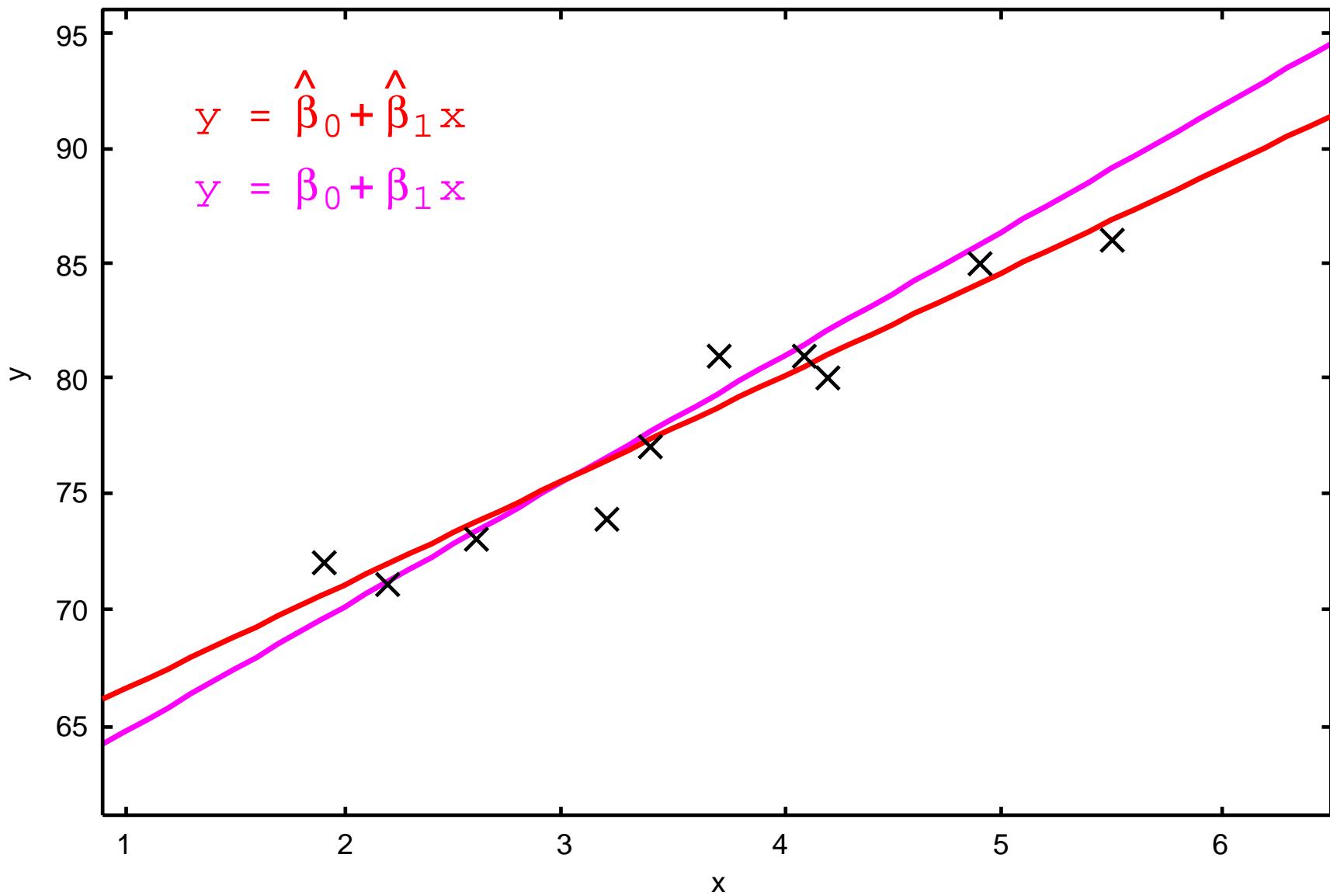
データの散布図



単回帰モデル $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$



直線の推定



5.2 説明変数が2個の場合の解析方法

データ (x_{i1}, x_{i2}, y_i)

No.	説明変数		目的変数 y
	x_1	x_2	
1	x_{11}	x_{12}	y_1
2	x_{21}	x_{22}	y_2
.	.	.	.
.	.	.	.
i	x_{i1}	x_{i2}	y_i
.	.	.	.
.	.	.	.
n	x_{n1}	x_{n2}	y_n
平均	\bar{x}_1	\bar{x}_2	\bar{y}

重回帰モデル

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \varepsilon_i, \quad \varepsilon_i \sim N(0, 1) \quad (5.3)$$

(ε_i は互いに独立に $N(0, \sigma^2)$ に従う。)

(1) 最小2乗法による回帰式の推定

予測値と誤差

$$\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_{i1} + \hat{\beta}_2 x_{i2} \quad (5.4)$$

$$e_i = y_i - \hat{y}_i = y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_{i1} + \hat{\beta}_2 x_{i2}) \quad (5.5)$$

残差平方和

$$S_e = \sum_{i=1}^n e_i^2 = \sum_{i=1}^n \{y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_{i1} + \hat{\beta}_2 x_{i2})\}^2 \quad (5.6)$$

单回帰分析のときと同様に, S_e を最小にするような,
 $\hat{\beta}_0, \hat{\beta}_1, \hat{\beta}_2$ を $\beta_0, \beta_1, \beta_2$ の推定値とする.

S_e を偏微分

(ホワイトボードで)

$$\frac{\partial S_e}{\partial \hat{\beta}_0} = -2 \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_{i1} - \hat{\beta}_2 x_{i2}) = 0 \quad (5.7)$$

$$\frac{\partial S_e}{\partial \hat{\beta}_1} = -2 \sum_{i=1}^n x_{i1} (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_{i1} - \hat{\beta}_2 x_{i2}) = 0 \quad (5.8)$$

$$\frac{\partial S_e}{\partial \hat{\beta}_2} = -2 \sum_{i=1}^n x_{i2} (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_{i1} - \hat{\beta}_2 x_{i2}) = 0 \quad (5.9)$$

(5.7) ~ (5.8) を整理(ホワイトボードで)

$$n\hat{\beta}_0 + \hat{\beta}_1 \sum_{i=1}^n x_{i1} + \hat{\beta}_2 \sum_{i=1}^n x_{i2} = \sum_{i=1}^n y_i \quad (5.10)$$

$$\hat{\beta}_0 \sum_{i=1}^n x_{i1} + \hat{\beta}_1 \sum_{i=1}^n x_{i1}^2 + \hat{\beta}_2 \sum_{i=1}^n x_{i1}x_{i2} = \sum_{i=1}^n x_{i1}y_i \quad (5.11)$$

$$\hat{\beta}_0 \sum_{i=1}^n x_{i2} + \hat{\beta}_1 \sum_{i=1}^n x_{i2}x_{i1} + \hat{\beta}_2 \sum_{i=1}^n x_{i2}^2 = \sum_{i=1}^n x_{i2}y_i \quad (5.12)$$

⇒ 正規方程式

$\hat{\beta}_0$ の計算

(5.10) 式より,

$$\begin{aligned}\hat{\beta}_0 &= \frac{1}{n} \sum_{i=1}^n y_i - \hat{\beta}_1 \frac{1}{n} \sum_{i=1}^n x_{i1} + \hat{\beta}_2 \frac{1}{n} \sum_{i=1}^n x_{i2} \\ &= \bar{y} - \hat{\beta}_1 \bar{x}_1 - \hat{\beta}_2 \bar{x}_2.\end{aligned}\tag{5.13}$$

$$\bar{y} = \hat{\beta}_0 + \hat{\beta}_1 \bar{x}_1 + \hat{\beta}_2 \bar{x}_2\tag{5.14}$$

推定された重回帰式は点 $[\bar{x}_1, \bar{x}_2, \bar{y}]$ を通る.

(5.13) を (5.11) に代入

$$\left(\frac{1}{n} \sum_{i=1}^n y_i - \hat{\beta}_1 \frac{1}{n} \sum_{i=1}^n x_{i1} + \hat{\beta}_2 \frac{1}{n} \sum_{i=1}^n x_{i2} \right) \sum_{i=1}^n x_{i1} \\ + \hat{\beta}_1 \sum_{i=1}^n x_{i1}^2 + \hat{\beta}_2 \sum_{i=1}^n x_{i1} x_{i2} = \sum_{i=1}^n x_{i1} y_i \quad (5.15)$$

(5.13) を (5.12) に代入

$$\left(\frac{1}{n} \sum_{i=1}^n y_i - \hat{\beta}_1 \frac{1}{n} \sum_{i=1}^n x_{i1} + \hat{\beta}_2 \frac{1}{n} \sum_{i=1}^n x_{i2} \right) \sum_{i=1}^n x_{i2} \\ + \hat{\beta}_1 \sum_{i=1}^n x_{i2} x_{i1} + \hat{\beta}_2 \sum_{i=1}^n x_{i2}^2 = \sum_{i=1}^n x_{i2} y_i \quad (5.16)$$

(5.15) を整理(ホワイトボード)

$$\begin{aligned} & \hat{\beta}_1 \left(\sum_{i=1}^n x_{i1}^2 - \frac{1}{n} \left(\sum_{i=1}^n x_{i1} \right)^2 \right) \\ & + \hat{\beta}_2 \left(\sum_{i=1}^n x_{i1} x_{i2} - \frac{1}{n} \sum_{i=1}^n x_{i2} \sum_{i=1}^n x_{i1} \right) \quad (5.17) \\ & = \sum_{i=1}^n x_{i1} y_i - \frac{1}{n} \sum_{i=1}^n x_{i1} \sum_{i=1}^n y_i \end{aligned}$$

(5.16) を整理(ホワイトボード)

$$\begin{aligned} & \hat{\beta}_1 \left(\sum_{i=1}^n x_{i1}x_{i2} - \frac{1}{n} \sum_{i=1}^n x_{i2} \sum_{i=1}^n x_{i1} \right) \\ & + \hat{\beta}_2 \left(\sum_{i=1}^n x_{i2}^2 - \frac{1}{n} \left(\sum_{i=1}^n x_{i2} \right)^2 \right) \\ & = \sum_{i=1}^n x_{i2}y_i - \frac{1}{n} \sum_{i=1}^n x_{i2} \sum_{i=1}^n y_i \end{aligned} \tag{5.18}$$

平方和 S_{11}, S_{22}

$$\begin{aligned} S_{11} &= \sum_{i=1}^n (x_{i1} - \bar{x}_1)^2 \\ &= \sum_{i=1}^n x_{i1}^2 - \frac{1}{n} \left(\sum_{i=1}^n x_{i1} \right)^2 \end{aligned} \tag{5.19}$$

$$\begin{aligned} S_{22} &= \sum_{i=1}^n (x_{i2} - \bar{x}_2)^2 \\ &= \sum_{i=1}^n x_{i2}^2 - \frac{1}{n} \left(\sum_{i=1}^n x_{i2} \right)^2 \end{aligned} \tag{5.20}$$

偏差積和 S_{12}

$$\begin{aligned} S_{12} &= \sum_{i=1}^n (x_{i1} - \bar{x}_1)(x_{i2} - \bar{x}_2) \\ &= \sum_{i=1}^n x_{i1}x_{i2} - \frac{1}{n} \sum_{i=1}^n x_{i2} \sum_{i=1}^n x_{i1} \end{aligned} \tag{5.21}$$

$$S_{11} \hat{\beta}_1 + S_{12} \hat{\beta}_2 = S_{1y} \quad (5.25)$$

$$S_{12} \hat{\beta}_1 + S_{22} \hat{\beta}_2 = S_{2y} \quad (5.26)$$

行列を使って書きなおすと、

$$\begin{bmatrix} S_{11} & S_{12} \\ S_{12} & S_{22} \end{bmatrix} \begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{bmatrix} = \begin{bmatrix} S_{1y} \\ S_{2y} \end{bmatrix} \quad (5.27)$$

と得る。

偏回帰係数

もし, $\det \begin{bmatrix} S_{11} & S_{12} \\ S_{12} & S_{22} \end{bmatrix} \neq 0$ ならば, この行列は逆行列を持ち,

$$\begin{aligned}\begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{bmatrix} &= \begin{bmatrix} S_{11} & S_{12} \\ S_{12} & S_{22} \end{bmatrix}^{-1} \begin{bmatrix} S_{1y} \\ S_{2y} \end{bmatrix} \\ &= \frac{1}{S_{11}S_{22}-S_{12}^2} \begin{bmatrix} S_{22} & -S_{12} \\ -S_{12} & S_{11} \end{bmatrix} \begin{bmatrix} S_{1y} \\ S_{2y} \end{bmatrix} \\ &= \frac{1}{S_{11}S_{22}-S_{12}^2} \begin{bmatrix} S_{22}S_{1y} - S_{12}S_{2y} \\ -S_{12}S_{1y} + S_{11}S_{2y} \end{bmatrix}. \tag{5.28}\end{aligned}$$

多重共線性

$\det \begin{bmatrix} S_{11} & S_{12} \\ S_{12} & S_{22} \end{bmatrix} = 0$ の場合, データに「多重共線性が存在する」という. このとき,

$$\begin{aligned} S_{11}S_{22} - S_{12}^2 = 0 &\Leftrightarrow \frac{S_{12}^2}{S_{11}S_{22}} = 1 \\ &\Leftrightarrow r_{x_1x_2}^2 = \left\{ \frac{S_{12}}{\sqrt{S_{11}S_{22}}} \right\}^2 = 1 \\ &\Leftrightarrow r_{x_1x_2} = \pm 1. \end{aligned} \tag{5.31}$$

多重共線性(続き)

即ち, 点 (x_{i1}, x_{i2}) ($i = 1, \dots, n$) の全てが一直線上に並んでいる状態である (p. 13 を見よ).

この場合, x_1 と x_2 のどちらか一方が分かれれば, もう一方が決まってしまうので, y を説明する変数は一つでいいということになる.

$r_{x_1 x_2}$ がちょうど 1 や -1 でなくとも, 1 か -1 に近い値の場合も同様のことが言えるし, (5.28) の計算が精度上まずくなることもある.

例題 1(時間があれば)